

Contents lists available at ScienceDirect

# Journal of Economic Behavior & Organization

journal homepage: www.elsevier.com/locate/jebo



# Distinguishing trust from risk: An anatomy of the investment game Daniel Houser<sup>a</sup>, Daniel Schunk<sup>b,\*</sup>, Joachim Winter<sup>c</sup>

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#### ARTICLE INFO

Article history:
Received 3 March 2009
Received in revised form
27 December 2009
Accepted 14 January 2010
Available online 1 February 2010

JEL classification:

C7

C9 D8

Keywords: Investment game Trust game Risk attitude Measurement

#### ABSTRACT

The role of trust in promoting economic activity and societal development has received considerable academic attention by social scientists. A popular way to measure trust at the individual level is the so-called "investment game" (Berg et al., 1995). It has been widely noted, however, that risk attitudes can also affect decisions in this game, and thus in principle confound inferences about trust. We provide novel evidence shedding light on the role of risk attitudes for trusting decisions. To the best of our knowledge our data are the first rigorous evidence that (i) aggregate investment distributions differ significantly between trust and risk environments, and (ii) risk attitudes predict individual investment decisions in risk games but not in the corresponding trust games. Our results are convergent evidence that trust decisions are not tightly connected to a person's risk attitudes, and they lend support to the "trust" interpretation of decisions in investment games.

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# 1. Introduction

The effect of trust on economic activity and development has received considerable interest in recent economic research (e.g., Guiso et al., 2006, 2008; Knack and Keefer, 1997; La Porta et al., 1997; Sapienza et al., 2007; Zak and Knack, 2001). In order to understand the role of trust as a determinant of economic activity, economists have begun to investigate empirically how trust affects the individual decisions of, and the interactions among, economic agents. This research has led many to question the nature of trust, and in particular to question the extent to which trusting decisions are connected to risk attitudes. To take one step towards addressing this issue, we here investigate whether trusting decisions in the widely used investment game (Berg et al., 1995) can be explained by a person's risk attitude.

Trusting decisions occur in environments of strategic uncertainty, where another person's decision affects one's own outcome (e.g., principle–agent relationships). Risky decisions occur when the environment includes state uncertainty (e.g., the outcome of the toss of fair dice.) Despite the conceptual distinction, scholars from various disciplines have argued that trust and risk are constructs that may be closely related in personal exchange contexts (see, e.g., Ben-Ner and Putterman, 2001; Hardin, 2002; Cook and Cooper, 2003). Knowing whether trust can be predicted by risk attitudes is important: if trusting is a risky decision, then policies to promote trust might best focus on creating rules that, for example, promote

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**Table 1** Summary of treatments.

Treatment	Trustor (Player 1)	Trustee (Player 2)	Receiver	Info about return distribution shown
Trust-1 (T1)	Human	Human	Human	No
Trust-2 (T2)	Human	Human	Human	Yes
Risk-1 (R1)	Human	Computer	-	Yes
Risk-2 (R2)	Human	Computer	Human	Yes

transparency and encourage peer-to-peer punishment of trust-violations. In contrast, if trust is not about risk, then such policies might be ineffective in promoting economic exchange.<sup>1</sup>

Experiments in the lab and field frequently use the "investment game" (Berg et al., 1995, henceforth BDM) to study trust.<sup>2</sup> However, several authors (Fehr, 2009; Karlan, 2005; Kosfeld et al., 2005; Sapienza et al., 2007) point out that in the investment game, inferences regarding trust may be confounded by individual attitudes towards risk. In this paper, we offer a novel approach to address rigorously the question of whether decisions in the BDM investment game can be predicted by a person's risk attitudes.

Discovering how risk attitudes influence trusting decisions in the BDM investment game is challenging for several reasons. One is that trust involves imperfect information over the likelihood of another person's decisions (strategic uncertainty), while risk is associated with perfect information over the likelihood of outcomes that often do not involve another person (state uncertainty). The fact that trust and risk games typically differ in multiple dimensions can make it difficult to connect trusting decision to attitudes towards risk (see, e.g., Bohnet and Zeckhauser, 2004, for additional discussion as well as a design that takes a first step towards addressing this issue).

A second challenge is that compelling inferences on this topic are not available from analyses based only on distributions of decisions between games. The reason is that aggregate distributions of decisions might be the same in trust and risk environments, and yet at the individual level decisions under risk might be unconnected to trusting decisions. This possibility might help to reconcile the conflicting results of, for example, Bohnet and Zeckhauser (2004) and Kosfeld et al. (2005). Both studies report results based only on distributions of decisions between games. The former find decisions differ between trust and risk environments, while the latter do not.

Collecting individual-level data on risk attitudes addresses both of these challenges. In particular, doing so allows one to move beyond aggregate analyses, and adds credibility to inferences regarding causal relationships between risk attitudes and trusting decisions. Even with such data in hand, inference can be subtle. For example, Eckel and Wilson (2004) investigate how behavior in two-person sequential, binary trust games correlates with a variety of behavioral and survey-based risk measures. Among their risk measures is a binary "risk" game similar to their binary trust game, as well as a Holt and Laury (2002, henceforth HL) measure of risk attitudes. They find that decisions in neither the binary risk game nor the HL game predict decisions to trust. However, they also find that HL measures are unable to predict decisions in their risk game. The inability of HL risk measures to predict decisions in the baseline risk environment leaves it unclear how to interpret the failure of those same risk attitudes to predict trust decisions in a similar environment.<sup>3</sup>

Earlier studies comparing decisions between trust and risk environments, such as those discussed above, have taken important steps in distinguishing roles of trust and risk in decision making. Our investigation contributes to this literature by reporting data from a new design that, to the best of our knowledge, is the first to address rigorously each of the challenges discussed above. Our approach involves combining measures of individual risk attitudes with individual decisions in investment games that do and do not include a "trust" component. More specifically, our procedure involves conducting two "trust" treatments and two "risk" treatments with the investment game. Each participant played exactly one of these games, and the games were run as separate treatments. A summary of our treatments is provided in Table 1.

In each trust treatment the trustee is a human, while in each risk treatment the trustee's decision is determined by a computer. In one risk treatment there is no human trustee so that the investor faced a standard individual decision problem

<sup>&</sup>lt;sup>1</sup> For example, some authors suggest betrayal aversion might discourage trusting decisions (see, e.g., Bohnet and Zeckhauser, 2004), and having legal recourse might do little to ameliorate betrayal's sting. Thus, knowing how to design institutions to promote trust, growth, and social development seems to require knowing whether and how trust is related to risk attitudes.

<sup>&</sup>lt;sup>2</sup> Earlier "gift exchange" experiments include Camerer and Weigelt (1988) and Fehr et al. (1993). An alternative approach to analyzing the role of trust on economic behavior is the use of data from surveys that contain direct questions on trust, often based on a question that was pioneered in the World Values Survey. Some studies also implement trust experiments within field surveys. Examples of survey-based studies are Glaeser et al. (2000), Fehr et al. (2002), Guiso et al. (2008), Ermisch et al. (2009), and Sapienza et al. (2007). Trust games are also implemented in neuroeconomic studies (e.g. Knoch et al., 2009).

<sup>&</sup>lt;sup>3</sup> A related point can be made of Glaeser et al. (2000), who investigate (but find only limited evidence for) correlations between survey-based attitudinal measures of trust and decisions in trust games. Fehr et al. (2002), in contrast, find that survey-based measures of trust are indeed correlated with the sender's behavior in a trust game. However, neither study addresses risk: These authors do not investigate correlations between their survey responses and decisions in otherwise identical trust and risk games. Similarly, Schechter (2007) compares risk and trust games without providing evidence that risk attitudes explain behavior in former.

under risk; while in a second risk treatment a computer made a decision for a passive human "trustee." This controls for prosocial impulses that might drive decisions in the trust game environment.<sup>4</sup>

Moreover, one trust treatment is a standard trust game in which investors receive no information about trustee behavior, while in the other trust treatment investors received information about "typical" returns (we followed the BDM social history treatment.) Both risk treatments included information on the return distribution used by the computer, which was again taken from BDM. This approach allows us to account for information differences between trust and risk environments.

Finally, in all four treatments we measure each subject's risk attitudes using the HL risk instrument.<sup>5</sup> As we noted above, these data are useful only to the extent that HL measures predict decisions in at least one of our treatments. Our key hypothesis is that the HL risk attitudes predict decisions in risk treatments, but not trust treatments. We investigate this hypothesis in two ways. First, because investors can "opt out" of playing the game by simply choosing not to invest, we ask whether risk attitudes predict this opt out decision. Second, among those who choose to invest, we investigate whether the investment amount is predicted by risk attitudes.

We find no connection between risk attitudes and the decision to opt out (invest zero) in either the trust or risk treatments. Conditional on choosing to invest a positive amount, however, we obtain clear evidence that risk attitudes predict decisions in risk treatments. We are unable to discover a predictive relationship between risk attitudes and decisions in trust treatments. This finding does not necessarily imply that risk attitudes are unimportant to trusting decisions, but it does suggest that, to the extent that risk attitudes do modulate trusting decisions, the mechanism remains to be discovered. The findings thus suggest a fundamental distinction between risks constituted by non-social factors and risks based on interpersonal interactions. Methodologically, our findings suggest that measures of trust derived from the investment game are not predicted by measures of risk derived from a price list procedure.

Substantively, our results offer rigorous support for the view that motives for trust are not tightly connected to risk attitudes. This leaves open the possibility that emotional factors such as betrayal aversion (Bohnet et al., 2008; de Quervain et al., 2004; Aimone and Houser, 2008) play an important role in mediating trusting decisions. Second, it suggests that the power of trust to explain various economic outcomes – including stock market participation, cash holdings, credit card usage, and foreign direct investments – is not entirely due to a close connection between trusting and risky decisions (e.g. Guiso et al., 2004, 2008).

We proceed by presenting the design of our experiment (in Section 2). We report our results in Section 3, and Section 4 concludes.

# 2. Experiment design

#### 2.1. Procedures

The experiments included a total of 291 subjects and were conducted in the experimental laboratory of the Sonder-forschungsbereich 504, a research center at the University of Mannheim. Initially 117 subjects participated in two treatments in November 2005 (denoted as Trust-1 and Risk-1 in the discussion below), another 96 subjects participated in a second trust treatment in April 2007 (denoted as Trust-2), and another 78 subjects participated in a second risk treatment in December 2007 (denoted as Risk-2). All subjects were recruited from the general student population. The median age of the participants was 23 years, and 37% of the participants are female.

The experiment lasted between 22 and 35 min. Each treatment consists of a HL risk attitude elicitation task (henceforth, the "HL task") and a trust or risk game. Half of our subjects completed the HL task first, and the other half the trust or risk game first. We found no evidence of order effects.

# 2.2. Risk elicitation task and trust and risk games

We first describe the HL task we used to draw inferences regarding participants' degrees of risk aversion. The task is a replication of the price list procedure used by Holt and Laury (2002). It involves ten choices between the paired lotteries A and B described in Table 2. The consequences of lotteries A and B are the same in all 10 choice situations, and lottery B always has a higher variability than lottery A. However, the probabilities associated with the consequences of the lotteries change across the rows: while in the first row, the probability of the high payoff for both options is 10%, it increases to 100% in the last row. A very risk-seeking person should thus switch to option B early, and an extremely risk-averse person should switch over by decision 10 in the bottom row. Following Holt and Laury (2002), payoffs for each subject were determined by randomly implementing one of the ten lotteries and paying according to the subject's decision on that lottery.

Turn now to our trust and risk games. The *Trust-1* (or T1) treatments follow exactly the procedures of the standard BDM investment game. Participants are randomly and anonymously paired, and each is endowed with 10 experimental currency

<sup>&</sup>lt;sup>4</sup> A study of connections between pro-social preferences and trusting decisions is reported by Cox (2004). Our paper is not intended to investigate this

<sup>&</sup>lt;sup>5</sup> A variety of procedures are available to elicit risk attitudes, including Abdellaoui (2000), Hey and Orme (1994), Wakker and Deneffe (1996), and Charness and Gneezy (2010).

**Table 2**Lottery choices in the HL risk elicitation procedure.

Option A	Option B
10% of €2.00, 90% of €1.60	10% of €3.85, 90% of €0.10
20% of €2.00, 80% of €1.60	20% of €3.85, 80% of €0.10
30% of €2.00, 70% of €1.60	30% of €3.85, 70% of €0.10
40% of €2.00, 60% of €1.60	40% of €3.85, 60% of €0.10
50% of €2.00, 50% of €1.60	50% of €3.85, 50% of €0.10
60% of €2.00, 40% of €1.60	60% of €3.85, 40% of €0.10
70% of €2.00, 30% of €1.60	70% of €3.85, 30% of €0.10
80% of €2.00, 20% of €1.60	80% of €3.85, 20% of €0.10
90% of €2.00, 10% of €1.60	90% of €3.85, 10% of €0.10
100% of €2.00, 0% of €1.60	$100\% \text{ of } \in 3.85, 0\% \text{ of } \in 0.10$

units (ECU). Subjects exchange ECU for Euros at a rate of one Euro for two ECU at the experiment's conclusion. Each investor can send some, all or none of her endowment to her counterpart, the trustee. The experimenter triples the amount sent and provides that amount to the trustee, who can then return some, all, or none of the tripled amount to the investor. "Trust" is often measured as the amount investors send in this game, though we noted many have pointed out that risk attitudes might motivate decisions to "trust" in this environment.

The *Trust-2* (or T2) treatments were conducted in the same way as T1, except that investors were given the information provided by Berg et al. (1995) to their subjects in their social history treatment. Participants were told that this was a description about how people had made decisions in this game in the past, but were aware that it was not a guarantee of how decisions might be made in their session. This treatment is important because this information was always provided to subjects in our risk treatments. Thus, if HL measures fail to predict decisions in this trust treatment, but predict decisions in our risk treatments, it is not likely attributable to differences in information conditions between treatments.

The risk treatments are also modeled on the investment game, but vary from the trust treatments in that the return decision is determined by a computer. In both risk treatments investors are shown a graph describing the computers' true return distribution. As noted above, this distribution is taken from Berg et al. (1995), and investors are informed that the distribution is based on previous experiments with human subjects. The risk treatments vary in that Risk-1 (R1) is a pure individual decision problem – there is no other human involved in the game. It is important to know whether HL measures can predict decisions in this game, in order to be able to interpret any potential failure of HL to predict decisions in trust treatments.

Our final treatment, Risk-2 (R2), is the same except that it includes a passive human "trustee" whose payment is entirely determined by the investor's and computer's decisions. We ran this treatment to account for the presence of prosocial motivations for investment in the trust treatments. If HL predicts decisions in R2, but not the trust treatments, this suggests that differences do not stem from simply the presence of a person in the trust treatments, but can perhaps instead be traced to the need to trust that person.

Finally, note that all participants had full information about the game they were playing. In particular, all were aware whether the "trustee" decision was made by a human or computer.

# 2.3. Distinguishing decisions between trust and risk environments

It is worthwhile to reiterate that our design allows us to compare decisions made in a trust environment with those made in an otherwise identical risk environment. That is, our treatments are built to control systematically for differences between the two environments in prosocial impulses or information. The key remaining difference is that in the "trust" environment any backtransfer amount is determined by a human (an environment of strategic uncertainty), while in the risk environment backtransfers are determined by an automatic process not connected to a human participant (an environment of state uncertainty).

The investigation of decision patterns between our trust and risk environments informs us whether decisions differ between cases of strategic and state uncertainty. Further, by collecting individual data on risk attitudes, and then comparing those data to decisions made in our games, we can provide further evidence on the reason for any between-treatment differences. For example, finding that risk attitudes predict decisions in our risk treatments, but not our trust treatments, is evidence that subjects perceived the environment of state uncertainty as involving "risk", that they perceived the environment strategic uncertainty differently, and that they made decisions between those two environments in different ways.

<sup>&</sup>lt;sup>6</sup> Note that the decision to trust has a positive financial return whenever the trustee returns more than 33% of the investor's transfer amount.

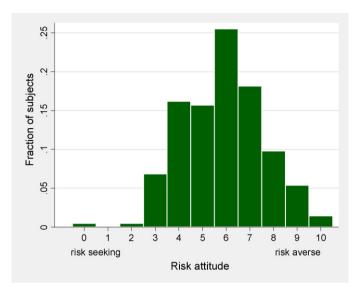


Fig. 1. Distribution of the HL risk measure (all subjects).

#### 3. Results

In this section we first describe our subjects' risk attitudes as elicited by the HL task. Then we show how the distributions of investor decisions differ between our trust and risk treatments. Finally, we provide evidence that risk attitudes predict decisions in risk games but do not predict decisions in trust games.

# 3.1. Subjects' risk attitudes

In total, we observe 204 subjects as senders. Their risk aversion is measured according to Holt and Laury (2002) as the (last) point where a subject switches from option A to option B.<sup>7</sup> Very risk-seeking subjects have a low switchpoint, and risk-averse subjects a high switchpoint. We find a mean switchpoint of 5.86, suggesting that subjects are risk averse on average. Similar to Holt and Laury we also find that more than two thirds of the subjects (76%) choose more than 4 safe choices, which is the predicted switchpoint of a risk-neutral subject. 76 (37%) of our subjects are female, and female participants are on average more risk averse (mean switchpoint: 6.07) than male participants (mean switchpoint: 5.75); although this difference is statistically insignificant. Fig. 1 displays the distribution of risk attitudes for our 204 participants. The distribution of risk attitudes does not differ across our four treatments (Kolmogorov–Smirnov-tests, all *p*-values >0.2).

# 3.2. Investment distributions in risk and trust games

Fig. 2 shows histograms of investment decisions in the two trust conditions T1 and T2 (n = 48 for both treatments) and the two risk treatments R1 and R2 (n = 69 and n = 39), respectively. Turning first to the trust treatments, their respective histograms have in common significant mass at the right tail and relatively flat middle sections. Although it is apparent that a greater fraction chose not to invest in T1,<sup>8</sup> there is nevertheless only weak statistical evidence for differences between their overall distributions (p = 0.12, Kolmogorov–Smirnov test); their means (p = 0.10); their medians (p = 0.11) or their variances (p = 0.07).<sup>9</sup> This finding suggests that the prior that we induce in T2, R1, and R2 by showing information about behavior in the BDM does not differ much from the prior that subjects had anyway.

The risk conditions R1 and R2 are also highly visually similar, and we find no evidence for differences along any of the dimensions described above (p > 0.25 for the distribution test, means test, medians test and variance test). An important first result emerges here: in our data, the presence of a human as receiver (R2) does not change behavior compared to the same game absent a human receiver (R1).

<sup>&</sup>lt;sup>7</sup> The vast majority of our subjects (95%) switched only once. As in Holt and Laury (2002), we find that the analysis reported in this paper changes very little if we drop subjects who switch from B back to A.

<sup>&</sup>lt;sup>8</sup> Recall that T1 is the only case where subjects have no information regarding return distributions. Our results might suggest that such ambiguity can have a detrimental effect on participation rates in these sorts of games.

<sup>&</sup>lt;sup>9</sup> Moreover, any suggestion of differences is driven only by differences in participation rates (zero investments). As we discuss below, when considering only positive investment amounts all hints of differences vanish.

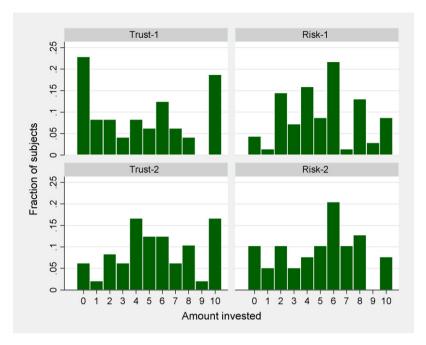


Fig. 2. Distribution of investment decisions in the human and computer treatments.

Our second finding, which we develop in further detail in the remaining parts of the paper, is inspired by the observation that both risk games (R1 and R2) differ from the two trust games (T1 and T2). The risk conditions follow a pattern with substantially more mass concentrated near the center, and much less at the tails, than is found in the trust conditions (T1 and T2). As a result the variance of the trust distributions is greater than the risk distributions. Indeed, the difference between the variances in treatments T1 and R1 is statistically significant (p = 0.01, variance ratio test.)

A source of differences between the trust and risk investment distributions could be that risk attitudes affect risky decisions but not trusting decisions. Another is that risk attitudes systematically affect decisions in both games but in different ways, or that they affect decisions in both games in the same way. The advantage to collecting the HL risk preference data is that it allows us to provide evidence on these possibilities. In particular, if decisions in both the trust and risk treatments are systematically mediated by risk attitudes then the HL data would be expected to predict either both (if HL in fact predicts risky decisions in this environment) or neither (if HL measures do not for some reason predict behavior in this environment.) On the other hand, the first possibility would be supported if risk preferences were found to be connected to risky decisions, but not trust decisions.

We are not able to detect simple linear relationships between subjects' HL risk attitudes and their decisions in trust or risk games (the simple correlation coefficients are small and insignificantly different from zero). Instead, we adopt a type-classification approach to discerning relationships between risk attitudes and decisions in our games. Type-classification has proven to be a powerful procedure to discerning economic relationships that are not apparent with alternative analyses (see, e.g., Houser et al., 2004). The next section turns to this approach.

#### 3.3. Investment and risk preference in trust and risk games

To explore the effect of risk attitudes on decisions in our games we classify subjects according to their degree of risk aversion. There are very many ways that one might do this, <sup>11</sup> so our approach is to draw our classifications directly from Holt and Laury (2002). In particular, HL draws attention to two particularly interesting patterns: four safe followed by six risky (switch at the fifth gamble), and six safe followed by four risky (switch at the seventh gamble). Theory predicts the former pattern for risk-neutral subjects, and the latter pattern for subjects who are risk averse in a way that is consistent with what has been found in various econometric analyses of auction data (Holt and Laury, 2002, p. 1646). We follow this

<sup>&</sup>lt;sup>10</sup> Clustering at boundaries in trust games has been reported by others, e.g., Kosfeld et al. (2005). See also a survey by Camerer (2003). Note that this finding is robust to the participation decision, i.e. it also holds if we drop zero investments from the sample.

<sup>&</sup>lt;sup>11</sup> For example, there are 121 unique ways to assign three risk-preference types based on HL switch patterns.

approach directly, and classify subjects into three risk preference categories: those who switch at or before the fourth HL gamble (risk seeking); those who switch at the fifth or sixth HL gamble (risk neutral); and those who switch at or after the seventh HL gamble (risk averse).

Table 3 details the relationship between subjects' risk types and their investment decisions in each of our four treatments. For each treatment and for each risk "type," we report various investment statistics: mean, median, high, low, and whether a subject invested more than one third of his/her endowment.<sup>12</sup> Consider first the trust treatments. Mean invested amounts do not vary systematically with risk attitudes in either treatment, and we are unable to obtain formal evidence that median contributions monotonically increase with risk attitudes (p=0.13 and p=0.82, Jonckheere tests, <sup>13</sup> T1 and T2, respectively). Moreover, about one-half of each risk type invest four or more in T1, while about 4/5 do so in T2. In pairwise comparisons, the fraction of subjects investing four or more is statistically identical between types within T1 and within T2 (the minimum p-value is 0.19 among the three pairwise, two-sided t-tests in each treatment).

Turn next to the risk treatments, where the data patterns are different. Note first that both mean and median contributions increase as subjects' risk preferences change. Means range from 4.7 to 5 to 5.9 as risk type increases from risk averse to risk seeking in the R1 treatment, and 4.4 to 5.0 to 6.5 in R2. Medians also vary monotonically in both treatments, and these ordered relationships are statistically significant (p < 0.05 in both cases, Jonckheere tests). Strikingly, 90% and 100% of risk-seeking types invest four or more in R1 and R2, respectively, while this is true for only about 60% of these treatments' risk-averse participants. Again, these differences are statistically significant (p < 0.01 for R1 and p = 0.03 for R2).

### 3.4. Participation and risk preference in risk and trust games

Next it is worthwhile to briefly consider what can be discovered about the impact of risk attitudes on the participation decision. As can be seen from Table 3, we have few non-participation decisions (21 in total; 14 in the trust treatments and 7 in the risk treatments). Nevertheless, something can be learned by aggregating the human and computer treatments. Doing so reveals that in the risk treatments 9.1% of risk-averse participants chose not to participate, 6.1% of risk-neutral subjects and just 3.8% of risk-seeking subjects. The differences are not significant, but the direction is as one might expect. On the other hand, in the trust treatments the corresponding zero-investment frequencies are 22.5% (risk averse), 11.67% (risk neutral) and 17.4% (risk seeking). These numbers are greater than in the risk games, 14 but display no systematic relationship to risk preference.

To provide additional evidence on the participation decision we conducted a probit analysis on the pooled data set. Note that pooling the data is statistically valid.<sup>15</sup> The results of this analysis are reported in Table 4. The dependent variable is whether to invest or not, and the independent variables include interactions of a person's risk type and the trust or risk condition.<sup>16</sup> Also, we include in the analysis dummies for gender and economics major, as well as an indicator for age. As is clear from Table 4, our finding is that risk attitudes do not seem to have an impact on participation decisions. The coefficients are statistically insignificant and small in magnitude.

Next, we complement the analysis in Section 3.3 by investigating whether risk attitudes predict whether a person who contributes is likely to fall into the low or high contribution column. Our discussion and analysis of the data in Table 3 certainly suggest this should be the case for the risk treatments, but not the trust treatments. We executed the probit model described above using the group of subjects who contributed a positive amount, and with dependent variable the decision to make an investment of four or more.<sup>17</sup> Note that we continue to pool the trust and risk treatments for this analysis, and this is statistically valid (tests for pooling yield p = 0.54 and p = 0.48 in the risk and trust treatments, respectively, where the pooling test is analogous to that described in fn. 15).

The results of our analysis are described in Table 5, and are convergent with the patterns discovered in our analysis surrounding Table 3. In particular, subjects who were classified as "risk seeking" on the HL scale are statistically significantly more likely to contribute a high amount in the risk treatments. The effect is also economically large, increasing from a probability of roughly 0.75 for risk neutral or risk averse subjects to about 0.98 for subjects who are risk seeking. On the

 $<sup>^{\</sup>rm 12}\,$  Using a cut-point of 20% or 40% leaves our findings from this analysis unchanged.

<sup>&</sup>lt;sup>13</sup> Jonckheere (1954) develops a non-parametric test for ordered relationships. In the case of two samples it reduces to the Mann–Whitney test. In our case the null hypothesis is that there are no systematic relationships among the medians of the different types' investment distributions, against the alternative that the medians are ordered from risk-averse (lowest) to risk-seeking (highest).

<sup>&</sup>lt;sup>14</sup> A greater non-participation rate in investment games with human counterparts, as compared to games with computer counterparts, is also reported by Aimone and Houser (2008). Those authors attribute the difference to betrayal aversion.

<sup>&</sup>lt;sup>15</sup> We ran the probit regression model reported in Table 3 using full interactions of the three risk categories with the complete set of treatment dummies. Based on this regression we neither reject the hypothesis that the coefficients for the two computer treatments are jointly identical (p = 0.85), nor do we reject the hypothesis that the coefficients of the two human treatments are jointly identical (p = 0.38).

<sup>&</sup>lt;sup>16</sup> In our regressions, the baseline category is "Risk seeking  $\times$  computer".

<sup>&</sup>lt;sup>17</sup> One can also conduct a multivariate (or ordered) probit with categories: zero, one-three, and four or greater. Doing this leads to the same findings and does not shed any additional light on the issue at hand, but complicates the interpretation of the coefficient estimates. For ease of explication we report two probit analyses.

**Table 3**Risk types and investment decisions in the Trust-1, Risk-1, Trust-2, and Risk-2 treatment.

	Mean investment [ECU]	Median investment [ECU]	Lowest investment [ECU]	Highest investment [ECU]	Investment ≥4 ECU		0 < Investment < 4 ECU		Investment = 0 ECU	
					%	n	%	n	%	n
Trust-1 treatment										
Risk averse	3.6	3	0	10	44%	8	17%	3	39%	7
Risk neutral	5.1	5	0	10	67%	12	28%	5	5%	1
Risk seeking	4.7	5	0	10	58%	7	17%	2	25%	3
Total	4.4	4	0	10	56%	27	21%	10	23%	11
Risk-1 treatment										
Risk averse	4.7	4	0	10	63%	12	26%	5	11%	2
Risk neutral	5.0	5	1	10	67%	20	33%	10	0%	0
Risk seeking	5.9	6	0	10	90%	18	5%	1	5%	1
Total	5.2	5	0	10	72%	50	23%	16	5%	3
Trust-2 treatment										
Risk averse	5.6	5	1	10	75%	15	25%	5	0%	0
Risk neutral	5.1	5	0	10	76%	13	12%	2	12%	2
Risk seeking	6.1	6	0	10	82%	9	9%	1	9%	1
Total	5.5	5	0	10	77%	37	17%	8	6%	3
Risk-2 treatment										
Risk averse	4.4	4	0	10	57%	8	36%	5	7%	1
Risk neutral	5	6	0	10	68%	13	16%	3	16%	3
Risk seeking	6.5	6	4	10	100%	6	0%	0	0%	0
Total	5	6	0	10	69%	27	21%	8	10%	4

**Table 4**Probit regression of whether to invest or not.

	Coefficient	Std. Err.	Z	P >  z
Risk averse × computer	-0.1539	0.4258	-0.36	0.718
Risk seeking × computer	0.2471	0.5342	0.46	0.644
Risk averse × human	-0.5713	0.3755	-1.52	0.128
Risk neutral × human	-0.1089	0.4233	-0.26	0.797
Risk seeking × human	-0.5315	0.4275	-1.24	0.214
Female	0.2316	0.2714	0.85	0.393
Old	0.2875	0.2575	-1.12	0.264
Econ-major	-0.1433	0.2609	-0.55	0.583
Constant	1.3532	0.3485	3.88	0.000

*Note*: Number of observations = 204, pseudo  $R^2$  = 0.0632.

**Table 5**Probit regression of investing four or more.

	Coefficient	Std. Err.	Z	P >  z
Risk averse × computer	-0.0636	0.3173	-0.20	0.641
Risk seeking × computer	1.3221	0.5202	2.54	0.011
Risk averse × human	0.1077	0.3148	0.34	0.732
Risk neutral × human	0.2219	0.3227	0.69	0.492
Risk seeking × human	0.5875	0.4153	1.41	0.157
Female	-0.0029	0.2299	-0.01	0.990
Old	-0.3793	0.2287	-1.66	0.097
Econ-major	-0.1618	0.2223	-0.73	0.467
Constant	0.8158	0.2740	2.98	0.003

*Note*: Number of observations = 183, pseudo  $R^2$  = 0.0683.

other hand, this analysis is again unable to produce evidence that risk attitudes relate in a systematic way to trusting decisions.

# 4. Conclusion

This paper provided new evidence on the role of risk attitudes in trusting decisions in an investment game. We pointed out that addressing this issue rigorously has proven challenging, but that doing so is critical as a rapidly growing literature links measures of trust to important economic outcomes and provides insights on the design of institutions to facilitate economic growth and development.

Our approach was to examine behavior in investment games in four treatments: two "trust" treatments with human trustees, and two "risk" treatments where return decisions were made by computers. Like the risk treatments, in one trust treatment investors had information on previous human return decisions. Like the trust treatments, in one risk treatment a human received any money earned by the computer trustee. In addition, we elicited risk attitudes from each of our subjects using the Holt and Laury (2002) procedure. Thus, our design accounts for asymmetries between environments of strategic and state uncertainty, and also allows for analyses at the individual level.

We found that the aggregate distributions of investment decisions with computer counterparts (Risk-1 and Risk-2) differed substantially from the distribution with human counterparts (Trust-1 and Trust-2). The key difference is that decisions in the trust treatment were more likely to be at the extremes (zero trust or full trust), while in the risk treatments the most common decision was to risk about half the endowment. Consequently, investment decisions in trust environments exhibit greater dispersion than those in risk environments.

With respect to individuals' risk attitudes, we discovered that in risk environments mean and median investment amounts, as well as the probability of an investment of four or more, all increase systematically as participants become risk-seeking. In addition, we found that willingness to participate in risk games increases as preferences become more risk-tolerant. On the other hand, we were unable to discover systematic relationships between trusting decisions and risk attitudes. That we found risk attitudes not to predict trusting decisions does not necessarily imply that risk attitudes are unimportant to trust decisions. Our findings are also consistent with the possibility that risk attitudes interact with trusting decisions in a more complicated way than we have been able to investigate here, or in a way that exhibits substantial variability among people.

Finally, our results complement a broader literature providing evidence on the biological basis of trust and risk. This literature has found, for example, that trust is modulated by factors that do not affect risk attitudes (Kosfeld et al., 2005), and that trusting decisions are implemented differently by the brain than risky decisions (see, e.g., McCabe et al., 2001). Further research on the foundations of trusting behavior, and the potentially complex ways in which it is modulated by risk attitudes, would be fertile ground for future research.

# Acknowledgements

The authors are grateful to seminar participants at the University of Mannheim and the University of Zurich, as well as to participants of the 2007 Annual Conference of the German Economic Association (VfS) in Munich, the 2008 International Meeting on Experimental and Behavioral Economics (IMEBE) in Alicante, and the 2008 conference of the European Economic Association in Milano. This paper further benefited from comments by Charles Bellemare, Yan Chen, John Dickhaut, Ernst Fehr, Urs Fischbacher, Glenn Harrison, Benedikt Herrmann, Sabine Kroeger, Kevin McCabe, Mary Rigdon, and Klaus Schmidt. An earlier version was circulated under the title "Trust Games Measure Trust".

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